

## Convergence of products of independent random variables to the log-Poisson law

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Let us have a set of independent random values  $\{\xi_{nk}\}$  in each sequence, where  $n \in \mathbb{N}$ ,  $k = 1, 2, 3, \dots, k_n$  and  $k_n \rightarrow \infty$ , as  $n \rightarrow \infty$ . Suppose that the distribution of  $\xi_{nk}$  is defined:

$\xi_{nk}$	$-e$	$-1$	$0$	$1$	$e$
$P$	$P_{nk}^-$	$q_{nk}^-$	$q_{nk}^0$	$q_{nk}^+$	$P_{nk}^+$

Let the log-Poisson distribution be  $\Pi_{\lambda, \mu, \alpha_0, \alpha_1}(u)$  with parameters  $\lambda \geq 0$ ,  $\mu \geq 0$ ,  $\lambda + \mu > 0$ ,  $0 < \alpha_0 \leq 1$ ,  $|\alpha_1| \leq \alpha_0 e^{-2\mu}$ .

**Theorem.** *The  $M$ -weak limit of random functions  $P(\prod_{k=1}^{k_n} \xi_{nk} < u)$  is  $\Pi_{\lambda, \mu, \alpha_0, \alpha_1}(u)$  if and only if the next five conditions are satisfied:*

$$\lim_{n \rightarrow \infty} \max_{1 \leq k \leq k_n} \frac{P_{nk}^+ + P_{nk}^-}{1 - q_{nk}^0} = 0, \quad (1)$$

$$\lim_{n \rightarrow \infty} \prod_{k=1}^{k_n} (1 - q_{nk}^0) = \alpha_0, \quad (2)$$

$$\lim_{n \rightarrow \infty} \prod_{k=1}^{k_n} (P_{nk}^+ + q_{nk}^+ - (P_{nk}^- + q_{nk}^-)) = \alpha_1, \quad (3)$$

$$\lim_{n \rightarrow \infty} \sum_{k=1}^{k_n} \frac{P_{nk}^+ + P_{nk}^-}{1 - q_{nk}^0} = \lambda + \mu, \quad (4)$$

$$\alpha_1 \left( \lim_{n \rightarrow \infty} \sum_{k=1}^{k_n} \frac{q_{nk}^+ P_{nk}^- - q_{nk}^- P_{nk}^+}{(q_{nk}^+ + P_{nk}^-)^2 - (q_{nk}^- + P_{nk}^+)^2} - \mu \right) = 0. \quad (5)$$

*Proof. Sufficiency.* Let us use Mellin's transformations. Because random variables  $\xi_{nk}$  are independent in each sequence, we have:

$$\omega_n^0(t) = \omega_{\prod_{k=1}^{k_n} \xi_{nk}}^0(t) = \prod_{k=1}^{k_n} (q_{nk}^+ + q_{nk}^- + (P_{nk}^+ + P_{nk}^-)e^{it}),$$

$$\omega_n^1(t) = \omega_{\prod_{k=1}^{k_n} \xi_{nk}}^1(t) = \prod_{k=1}^{k_n} (q_{nk}^+ - q_{nk}^- + (p_{nk}^+ - p_{nk}^-)e^{it}),$$

$$W_n(t) = \begin{pmatrix} \omega_n^0(t) & 0 \\ 0 & \omega_n^1(t) \end{pmatrix}.$$

The Mellin's transformation of the log-Poisson distribution  $\Pi_{\lambda, \mu, \alpha_0, \alpha_1}$  is

$$W(t) = \begin{pmatrix} \alpha_0 e^{(\lambda + \mu)(e^{it} - 1)} & 0 \\ 0 & \alpha_1 e^{(\lambda - \mu)(e^{it} - 1)} \end{pmatrix}.$$

Using common theory about M-weak limit of sequence of M-decreasing independent random variables (see, e.g., [2], thm. 6.1), we get that  $W_n(t) \rightarrow W(t)$  for every  $t \in \mathbb{R}$ . Hence

$$P\left(\prod_{k=1}^{k_n} \xi_{nk} < u\right) \xrightarrow{M} \Pi_{\lambda, \mu, \alpha_0, \alpha_1}(u),$$

and the sufficiency is proved.

*Necessity.* Let we have  $P(\prod_{k=1}^{k_n} \xi_{nk} < u) \xrightarrow{M} \Pi_{\lambda, \mu, \alpha_0, \alpha_1}(u)$ . Hence  $W_n(t) \rightarrow W(t)$ . When  $t = 0$ , we obtain  $\omega_n^0(t) \rightarrow \alpha_0, \omega_n^1(t) \rightarrow \alpha_1$ .

Conditions (2) and (3) are satisfied. On the other hand:

$$\lim_{n \rightarrow \infty} \prod_{k=1}^{k_n} \left( \frac{q_{nk}^+ + q_{nk}^-}{1 - q_{nk}^0} + \frac{p_{nk}^+ + p_{nk}^-}{1 - q_{nk}^0} e^{it} \right) = e^{(\lambda + \mu)(e^{it} - 1)} \quad (6)$$

for all  $t \in \mathbb{R}$ .

Let  $\{\eta_{nk}\}$ ,  $n \in \mathbb{N}$ ,  $k = 1, 2, 3, \dots, k_n$  and  $k_n \rightarrow \infty$ , when  $n \rightarrow \infty$  is a set of independent random variables in each sequence. Let us suppose that  $\eta_{nk}$  obtains values 0 or 1 with probabilities  $\frac{q_{nk}^+ + q_{nk}^-}{1 - q_{nk}^0}$  and  $\frac{p_{nk}^+ + p_{nk}^-}{1 - q_{nk}^0}$  respectively.

The characteristic function of random variable  $\sum_{k=1}^{k_n} \eta_{nk}$  is

$$W_n^*(t) = \prod_{k=1}^{k_n} \left( \frac{q_{nk}^+ + q_{nk}^-}{1 - q_{nk}^0} + \frac{p_{nk}^+ + p_{nk}^-}{1 - q_{nk}^0} e^{it} \right).$$

We have  $\lim_{n \rightarrow \infty} W_n^*(t) = e^{(\lambda + \mu)(e^{it} - 1)}$  by (6) for all  $t \in \mathbb{R}$ .

So we can say that

$$\lim_{n \rightarrow \infty} P\left(\sum_{k=1}^{k_n} \eta_{nk} = l\right) = \frac{(\lambda + \mu)^l}{l!} \text{ for all } l = 0, 1, 2, \dots$$

Hence

$$\left\{ \begin{aligned} \lim_{n \rightarrow \infty} P\left(\sum_{k=1}^{k_n} \eta_{nk} = 0\right) &= e^{-(\lambda + \mu)}, \\ \lim_{n \rightarrow \infty} P\left(\sum_{k=1}^{k_n} \eta_{nk} = 1\right) &= (\lambda + \mu)e^{-(\lambda + \mu)}, \\ \lim_{n \rightarrow \infty} P\left(\sum_{k=1}^{k_n} \eta_{nk} = 2\right) &= \frac{(\lambda + \mu)^2}{2!} e^{-(\lambda + \mu)}. \end{aligned} \right. \quad (7)$$

On the other hand,

$$\left\{ \begin{aligned} P\left(\sum_{k=1}^{k_n} \eta_{nk} = 0\right) &= \prod_{k=1}^{k_n} \frac{q_{nk}^+ + q_{nk}^-}{1 - q_{nk}^0}, \\ P\left(\sum_{k=1}^{k_n} \eta_{nk} = 1\right) &= \sum_{k=1}^{k_n} \frac{p_{nk}^+ + p_{nk}^-}{1 - q_{nk}^0} \prod_{\substack{k'=1 \\ k' \neq k}}^{k_n} \frac{q_{nk'}^+ + q_{nk'}^-}{1 - q_{nk'}^0}, \\ P\left(\sum_{k=1}^{k_n} \eta_{nk} = 2\right) &= \sum_{k'=1}^{k_n} \sum_{\substack{k''=1 \\ k'' > k'}}^{k_n} \frac{p_{nk'}^+ + p_{nk'}^-}{1 - q_{nk'}^0} \frac{p_{nk''}^+ + p_{nk''}^-}{1 - q_{nk''}^0} \prod_{\substack{k=1 \\ k \neq k', k''}}^{k_n} \frac{q_{nk}^+ + q_{nk}^-}{1 - q_{nk}^0}. \end{aligned} \right. \quad (8)$$

Because

$$\lim_{n \rightarrow \infty} \left( \left( \frac{P(\sum_{k=1}^{k_n} \eta_{nk} = 1)}{P(\sum_{k=1}^{k_n} \eta_{nk} = 0)} \right)^2 - 2 \frac{P(\sum_{k=1}^{k_n} \eta_{nk} = 2)}{P(\sum_{k=1}^{k_n} \eta_{nk} = 0)} \right) = 0,$$

we get

$$\begin{aligned} &\lim_{n \rightarrow \infty} \left( \left( \sum_{k=1}^{k_n} \frac{p_{nk}^+ + p_{nk}^-}{1 - q_{nk}^0} \frac{1 - q_{nk}^0}{q_{nk}^+ + q_{nk}^-} \right)^2 \right. \\ &\quad \left. - \sum_{k'=1}^{k_n} \sum_{\substack{k''=1 \\ k'' > k'}}^{k_n} \frac{p_{nk'}^+ + p_{nk'}^-}{1 - q_{nk'}^0} \frac{p_{nk''}^+ + p_{nk''}^-}{1 - q_{nk''}^0} \frac{1 - q_{nk'}^0}{q_{nk'}^+ + q_{nk'}^-} \frac{1 - q_{nk''}^0}{q_{nk''}^+ + q_{nk''}^-} \right) = 0 \end{aligned}$$

or

$$\sum_{k=1}^{k_n} \left( \frac{p_{nk}^+ + p_{nk}^-}{q_{nk}^+ + q_{nk}^-} \right)^2 = 0.$$

Hence the condition (1) is satisfied.

Let remember connections (7) and (8), then we get:

$$\lim_{n \rightarrow \infty} \frac{P(\sum_{k=1}^{k_n} \eta_{nk} = 1)}{P(\sum_{k=1}^{k_n} \eta_{nk} = 0)} = \lambda + \mu \quad \text{and} \quad \lim_{n \rightarrow \infty} \sum_{k=1}^{k_n} \frac{p_{nk}^+ + p_{nk}^-}{q_{nk}^+ + q_{nk}^-} = \lambda + \mu.$$

Since

$$\sum_{k=1}^{k_n} \frac{p_{nk}^+ + p_{nk}^-}{1 - q_{nk}^0} - \sum_{k=1}^{k_n} \frac{p_{nk}^+ + p_{nk}^-}{q_{nk}^+ + q_{nk}^-} \leq \max_{1 \leq k \leq k_n} \frac{p_{nk}^+ + p_{nk}^-}{1 - q_{nk}^0} \sum_{k=1}^{k_n} \frac{p_{nk}^+ + p_{nk}^-}{q_{nk}^+ + q_{nk}^-},$$

and condition (1) is proved, we get condition (4). It remains to prove condition (5). Because

$$\begin{aligned} \omega_n^1(t) &= \prod_{k=1}^{k_n} ((q_{nk}^+ + p_{nk}^+) - (q_{nk}^- + p_{nk}^-)) \\ &\times \exp \left\{ (e^{it} - 1) \left( \sum_{k=1}^{k_n} \frac{p_{nk}^+ + p_{nk}^-}{1 - q_{nk}^0} - 2 \sum_{k=1}^{k_n} \frac{q_{nk}^+ p_{nk}^- - q_{nk}^- p_{nk}^+}{(q_{nk}^+ + p_{nk}^+)^2 - (q_{nk}^- + p_{nk}^-)^2} \right) \right. \\ &\left. + B \max_{1 \leq k \leq k_n} \frac{p_{nk}^+ + p_{nk}^-}{1 - q_{nk}^0} \sum_{k=1}^{k_n} \frac{p_{nk}^+ + p_{nk}^-}{1 - q_{nk}^0} \right\} \rightarrow \alpha_1 e^{(\lambda - \mu)(e^{it} - 1)}, \end{aligned}$$

and using above proved limits (3) and (4), we get

$$\alpha_1 \left( \lim_{n \rightarrow \infty} \sum_{k=1}^{k_n} \frac{q_{nk}^+ p_{nk}^- - q_{nk}^- p_{nk}^+}{(q_{nk}^+ + p_{nk}^+)^2 - (q_{nk}^- + p_{nk}^-)^2} - \mu \right) = 0,$$

and condition (5) is satisfied. Theorem is proved.

## References

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- [2] G. Bareikis, J. Šiaulyš, *Nepriklausomų Atsitiktinių Dydžių Sandaugos* (1998).
- [3] Ю.Ю. Мачис, Об испытаниях Бернули с переменными вероятностями, *Liet. Matem. Rink.*, 19(4), 145–151 (1979).

## Nepriklausomų atsitiktinių dydžių sandaugų konvergavimas į logpuasono dėsnį

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Darbe nagrinėjamas serijų sekos nepriklausomų kiekvienoje serijoje atsitiktinių dydžių sandaugų M-silpnas konvergavimas į logpuasono dėsnį. Yra įrodoma teorema su penkiomis sąlygom.