

# Asymptotic expansions in the large deviation zones for the distribution function of sums of random variables in the series scheme

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Let  $\xi_j^{(n)}, j = 1, 2, \dots, n$ , be a triangular array of independent random variables (r.v.) with means  $\mathbf{E}\xi_j^{(n)} = 0$ , and variances  $\sigma_j^{(n)2} = \mathbf{E}\xi_j^{(n)2} > 0, j = \overline{1, n}$ .

Put

$$S_n = \sum_{j=1}^n \xi_j^{(n)}, \quad B_n^2 = \sum_{j=1}^n \sigma_j^{(n)2}, \quad Z_n = B_n^{-1} S_n, \quad (1)$$

$$F_{Z_n}(x) = \mathbf{P}(Z_n < x), \quad \Phi(x) = \int_{-\infty}^x \varphi(y) dy, \quad (2)$$

$$\varphi(x) = \frac{1}{\sqrt{2\pi}} \exp\left\{-\frac{1}{2}x^2\right\}, \quad \psi(x) = \varphi(x)(1 - \Phi(x))^{-1}. \quad (3)$$

The target of our work is to obtain large deviation theorems and exponential inequalities for the function  $\mathbf{P}(Z_n \geq x)$ . First we have to get the estimate of the  $k$ th-order cumulant  $\Gamma_k(Z_n)$  of the r.v.  $Z_n$ , defined by equality (1), where  $\Gamma_k(X) := \frac{1}{i^k} \frac{d^k}{dt^k} \ln f_X(t) \Big|_{t=0}$ ,  $k = 1, 2, \dots$ . Here  $f_X(t) = \mathbf{E} \exp\{itX\}$  is the characteristic function (ch.f.) of the r.v.  $X$ . In what follows  $a \vee b = \max\{a, b\}$ ,  $a \wedge b = \min\{a, b\}$ ,  $[m]$  is an integer part of number  $m$ , and  $\theta_i, i = 1, 2, \dots$  denote some quantity, not exceeding 1 in absolute value.

Let us say that r.v.'s  $\xi_j^{(n)}, j = 1, 2, \dots, n$ , are subject to condition  $(B_\gamma)$ , if there exist quantities  $\gamma \geq 0$  and  $K_j^{(n)} > 0$  such that

$$|\mathbf{E}(\xi_j^{(n)})^k| \leq (k!)^{1+\gamma} (K_j^{(n)})^{k-2} \sigma_j^{(n)2}, \quad k = 3, 4, \dots \quad (B_\gamma)$$

In order to obtain the asymptotic expansion of the distribution function  $F_{Z_n}(x)$  of the r.v.  $Z_n$  defined by equality (1), in large deviation zones, according to the General Lemma 1 [4], one must have the estimates of  $k$ th-order cumulants of the r.v.  $Z_n$ .

**PROPOSITION 1.** *If for r.v.  $\xi_j^{(n)}, j = \overline{1, n}$ , condition  $(B_\gamma)$  is fulfilled, then for the  $k$ th-order cumulant  $\Gamma_k(Z_n)$  of the r.v.  $Z_n$ , the estimate*

$$|\Gamma_k(Z_n)| \leq (k!)^{1+\gamma} \Delta_n^{2-k}, \quad k = 3, 4, \dots, \quad (4)$$

$$\Delta_n = K_n^{-1} B_n, \quad K_n := 2 \max_{1 \leq j \leq n} (K_j^{(n)} \vee \sigma_j^{(n)}), \tag{5}$$

holds.

*Proof of the Proposition 1* is obtain in [2]. Next, suppose that the r.v.  $\xi_j^{(n)}, j = \overline{1, n}$ , densities  $p_{\xi_j}(x)$  are bounded, i.e.,

$$\sup_x p_{\xi_j}(x) \leq C_j^{(n)} < \infty. \tag{D}$$

Let  $\xi_j^{(n)}(h)$  be a r.v. conjugate to the  $\xi_j^{(n)}, j = \overline{1, n}$ , with the distribution density  $p_{\xi_j(h)}(y) := e^{hy} p_{\xi_j}(y) \left( \int_{-\infty}^{\infty} e^{hy} p_{\xi_j}(y) dy \right)^{-1}$ . Denote  $S_n(h) = \sum_{j=1}^n \xi_j(h), Z_n(h) = B_n^{-1}(h)(S_n(h) - M_n(h))$ . Then, it is easy to check that

$$M_n(h) = \sum_{k=2}^{\infty} \frac{1}{(k-2)!} \Gamma_k(S_n) h^{k-1}, \quad B_n^2(h) = \sum_{k=2}^{\infty} \frac{1}{(k-2)!} \Gamma_k(S_n) h^{k-2}, \tag{6}$$

where the quantity  $h = h(x)$  is the solution to the equation  $x = M_n(h)/B_n$ . Since the r.v.'s  $\xi_j^{(n)}, j = \overline{1, n}$ , are independent, we have

$$f_{Z_n(h)}(t) = \mathbf{E} e^{itZ_n(h)} = \exp \left\{ -it \frac{M_n(h)}{B_n(h)} \right\} \prod_{j=1}^n f_{\xi_j(h)} \left( \frac{t}{B_n(h)} \right). \tag{7}$$

Let us denote  $R_{n,\gamma} = \int_{T_{n,\gamma}}^{T_n} |f_{n,\gamma}^*(t)| \frac{dt}{t}$ , where

$$f_{n,\gamma}^*(t) = \begin{cases} \sum_{k=0}^s \left(\frac{3}{2}\right)^k \frac{x^k}{k!} f_{Z_n}^{(k)}(t), & \gamma > 0, \\ f_{Z_n(h)}(t), & \gamma = 0, \end{cases} \tag{8}$$

$s = 2[(1/2)(\Delta_n^2/18)^{1/(1+2\gamma)}] - 2, f_{Z_n}^{(0)}(t) = f_{Z_n}(t) = \mathbf{E} \exp\{itZ_n\}$  and

$$\begin{aligned} T_{n,\gamma} &:= (3/8)(1 - x/\Delta_{n,\gamma})\Delta_{n,\gamma}, \quad 0 \leq x < \Delta_{n,\gamma}, \\ \Delta_{n,\gamma} &:= c_\gamma \Delta_n^{1/(1+2\gamma)}, \quad c_\gamma = (1/6)(\sqrt{2}/6)^{1/(1+2\gamma)}, \end{aligned} \tag{9}$$

**PROPOSITION 2.** *If for the r.v.  $\xi_j^{(n)}, j = \overline{1, n}$ , with  $\mathbf{E}\xi_j^{(n)} = 0$  and  $\sigma_j^{(n)2} > 0, j = \overline{1, n}$ , condition  $(B_\gamma)$  is fulfilled, then for each integer  $l, l \geq 3$  and  $T_n \geq T_{n,\gamma}$  in the interval  $0 \leq x < \Delta_{n,\gamma}$ , the relation*

$$\begin{aligned} \frac{1 - F_{Z_n}(x)}{1 - \Phi(x)} &= \exp \{L_{n,m}^*(x)\} \left\{ \frac{\psi(x)}{\psi(u_n(x))} \left( 1 + \sum_{\nu=1}^{l-3} P_{\nu,n}(u_n(x)) \right) + \theta_1(x+1) \right. \\ &\times \left. \left( \frac{c(l, \gamma, x)}{\Delta_n^{l-2}} + \frac{285\Delta_n \exp \left\{ - (1 - x/\Delta_{n,\gamma}) \sqrt{\Delta_{n,\gamma}} \right\}}{(1 - x/\Delta_{n,\gamma})} + \frac{6q}{T_n} + R_{n,\gamma} \right) \right\} \tag{10} \end{aligned}$$

holds. Here

$$L_{n,m}^*(x) = \sum_{k=3}^m \lambda_{k,n} x^k, \quad m = \begin{cases} (1/\gamma) + l - 1, & \gamma > 0, \\ \infty, & \gamma = 0. \end{cases} \tag{11}$$

The coefficients  $\lambda_{k,n}$  are found by formula (6) [4] and expressed by cumulants of r.v.  $Z_n$ . For this coefficients the estimate  $|\lambda_{k,n}| \leq (2/k)(16/\Delta_n)^{k-2}((k+1)!)^\gamma$ ,  $k = 3, 4, \dots$  holds. The quantity  $u_n(x) = x \left( 1 + \sum_{k=1}^{l-3} c_{k,n} x^k + \theta_2 c^*(l, \gamma) (x/\Delta_n)^{l-2} \right)$ , where  $c^*(l, \gamma) = 736l(l-1)(7/2)^{l-2}(l!)^\gamma$  and the coefficients  $c_{k,n}$  are expressed by the cumulants of the r.v.  $Z_n$  and found by formula (11) [4].

Polynomials  $P_{\nu,n}(u_n(x))$  are determined by relation (104) [4]. In particular,

$$\begin{aligned} P_{1,n}(u_n(x)) &= -\frac{1}{2}\Gamma_3(Z_n)\frac{1}{x} + \frac{3}{2}(2\Gamma_4(Z_n) - 3\Gamma_3^2(Z_n)) + \frac{1}{48}(72\Gamma_5(Z_n) \\ &\quad - 394\Gamma_3(Z_n)\Gamma_4(Z_n) + 267\Gamma_3^3(Z_n))x + \dots, \\ P_{2,n}(u_n(x)) &= \frac{1}{24}(3\Gamma_4(Z_n) - 5\Gamma_3^2(Z_n)) + \frac{1}{24}(3\Gamma_5(Z_n) \\ &\quad - 16\Gamma_3(Z_n)\Gamma_4(Z_n) + 15\Gamma_3^3(Z_n))x + \dots \end{aligned} \tag{12}$$

Full expressions of the quantities  $c(l, \gamma, x)$  and  $q$  are obtained by (9) and (58) [4], respectively.

Notice, that the proof of Proposition 2 can be easily derive on the basis of the proof of Proposition 2 [1], p. 294.

According to Proposition 2, we have to estimate the integral  $R_{n,\gamma}$  defined by relation (8) for  $\gamma = 0$  and  $\gamma > 0$ , respectively. The known estimates for characteristic functions and employment of condition (D) (V.Statulevičius [5]) will help us to achieve this aim. According to the definition of the quantity  $\Delta_{n,\gamma}$  by relation (9) in the case  $\gamma = 0$ , we have

$$T_{n,0} = (1/8)(1 - x/\Delta_{n,0})\Delta_{n,0}, \quad \Delta_{n,0} = c_0\Delta_n, \quad c_0 = (1/6)(\sqrt{2}/6). \tag{13}$$

**Theorem 1.** *If for the r.v.  $\xi_j^{(n)}$ , with  $\mathbf{E}\xi_j^{(n)} = 0$  and  $\sigma_j^{(n)2} > 0$ ,  $j = \overline{1, n}$ , conditions (B $_\gamma$ ) with  $\gamma = 0$  and (D) are fulfilled, then in the Cramer zone  $0 \leq x < \Delta_{n,0}$ , the asymptotic expansion (10) holds. Moreover,*

$$\begin{aligned} R_{n,0} &\leq 684 e^4 \pi \sqrt{2\pi} K_n \max_{1 \leq r_i \leq n} \prod_{i=1}^4 C_{r_i}^{1/4} \exp \left\{ -\frac{1}{24K_n^2} \sum_{k=1}^n \frac{1}{C_k^2} \right\} \\ &\quad + \frac{\pi^2}{2T_{n,0}} \exp \left\{ -\frac{1}{\pi^2} T_{n,0}^2 \right\}. \end{aligned} \tag{14}$$

The *proof* of Theorem 1 is obtained in [6].

**Theorem 2.** Let for the r.v.  $\xi_j^{(n)}$ , with  $\mathbf{E}\xi_j^{(n)} = 0$  and  $\overline{\sigma_j^{(n)}} > 0$ ,  $j = \overline{1, n}$ , conditions  $(B_\gamma)$  with  $\gamma > 0$ ,  $(D)$  and

$$\lim_{n \rightarrow \infty} \frac{1}{(1 \vee L_{1,n})\Delta_{n,\gamma}} \frac{1}{K_n^2} \sum_{j=1}^n \frac{1}{C_j^2} \geq d > 0 \tag{S}$$

be fulfilled. Then in the Linnik power zones  $0 \leq x < \Delta_{n,\gamma}$ , the asymptotic expansion (10) holds. Moreover,

$$R_{n,\gamma} \leq c_4(\gamma)(K_n/\Delta_n) \max_{1 \leq r_i \leq n} \prod_{i=1}^4 C_{r_i}^{1/4} \left\{ -\frac{5d}{2K_n^2} \sum_{k=1}^n \frac{1}{C_k^2} \right\} + T_{n,\gamma}^2 \exp \left\{ -\frac{1}{16\pi^2} T_{n,\gamma}^2 \right\} + \frac{5\pi^2 x^2}{8} T_{n,\gamma} \exp \left\{ -\frac{1}{\pi^2} T_{n,\gamma}^2 \right\}, \tag{15}$$

where  $T_{n,\gamma}$ ,  $K_n$  are defined by equalities (9), (5), and  $c_4(\gamma) = 192 e^2 \sqrt{2\pi} \cdot 24^\gamma$ .

To prove Theorem 2, it is necessary, by Proposition 2, to estimate the integral  $R_{n,\gamma}$ , which is defined by equality (8), with  $\gamma > 0$ . Follow to the proof of the Theorem 2 [1], we get inequality (15).

**Theorem 3.** For the r.v.  $Z_n$  defined by equality (1), exponential inequalities

$$\mathbf{P} \{ \pm Z_n \geq x \} = \begin{cases} \exp \left\{ -(8 \cdot 2^\gamma)^{-1} x^2 \right\}, & 0 \leq x \leq 2(2^\gamma \Delta_n)^{\frac{1}{1+2\gamma}}, \\ \exp \left\{ -\frac{1}{4} (x \Delta_n)^{\frac{1}{1+\gamma}} \right\}, & x \geq 2(2^\gamma \Delta_n)^{\frac{1}{1+2\gamma}}, \end{cases} \tag{16}$$

If inequality (4) is valid, then  $|\Gamma_k(Z_n)| \leq (k!/2)^{1+\gamma} 2^{1+\gamma} \Delta_n^{2-k}$ . The proof follows from this inequality and Lemma 2.4 [3].

EXAMPLE. Let  $\xi_j^{(n)} := a_{nj} \xi_j$ , where  $a_{nj}$  are nonnegative quantities, and  $\xi_j$  are nonidentically distributed r.v.'s with  $\mathbf{E}\xi_j = 0$  and  $\sigma_j^2 = \mathbf{E}\xi_j^2$ ,  $j = \overline{1, n}$ . Now  $S_n = \sum_{j=1}^n a_{nj} \xi_j$ ,  $Z_n = B_n^{-1} S_n$ , where  $B_n^2 = \sum_{j=1}^n a_{nj}^2 \sigma_j^2$ . Suppose that there exist quantities  $\gamma \geq 1$  and  $K > 0$  such that  $|\mathbf{E}\xi_j^k| \leq (k!)^{1+\gamma} K^{k-2} \sigma_j^2$ ,  $k = 3, 4, \dots$ . Further, suppose that for r.v.  $\xi_j$  there exists density  $p_{\xi_j}(x) \leq C_j < \infty$ . In this case, in condition  $(B_\gamma)$   $K_j^{(n)} = a_{nj} K$ , and  $\sigma_j^{(n)} = a_{nj} \sigma_j$ . Then the quantities defined by equalities (4) and  $(D)$  are

$$K_n = 2 \max_{1 \leq j \leq n} a_{nj} \{K \vee \sigma_j\}, \quad C_j^{(n)} = a_{nj}^{-1} C_j. \tag{17}$$

**References**

[1] L. Saulis, Asymptotic expansions of large deviations for sums of nonidentically distributed random variables, *Acta Appl. Math.*, **58**, 291–310 (1999).

- [2] D. Deltuvienė, L. Saulis, Asymptotic expansion of the distribution density function for the sum of random variables in the series scheme in large deviation zones, *Acta Appl. Math.*, **78**, 87–97 (2003).
- [3] L. Saulis, V. Statulevičius, *Limit Theorems for Large Deviations*, Kluwer Academic Publishers, Dordrecht, Boston, London (1991).
- [4] L. Saulis, Asymptotic expansions in large deviation zones for the distribution function of random variables with cumulants of regular growth, *Lith. Math. J.*, **36**, 365–392 (1996).
- [5] V. Statulevičius, Limit theorems for the density functions and asymptotic expansions for distributions of sums of independent random variables, *Theory Probab. Appl.*, **10**, 645–659 (1965).
- [6] D. Deltuvienė, Asymptotic expansion for the distribution function of series scheme of random variables in the large deviation Cramer zone, *Lith. Math. J.*, **42**, 691–696 (2002).
- [7] S.A. Book, A large deviation theorem, *Z. Wahrsch. Verw. Geb.*, **26** (1973).

## Atsitiktinių dydžių sumos serijų schemoje pasiskirstymo funkcijos asimptotiniai skleidiniai didžiųjų nuokrypių zonose

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Darbas skirtas nepriklausomų atsitiktinių dydžių (at.d.)  $\xi_j^{(n)}$ ,  $j = \overline{1, n}$  su vidurkais  $\mathbf{E}\xi_j^{(n)} = 0$  ir dispersijomis  $\sigma_j^{(n)2} = \mathbf{E}\xi_j^{(n)2} < \infty$  sumos pasiskirstymo funkcijos asimptotinių skleidinių gavimui didžiųjų nuokrypių Kramerio ir laipsninėse Liniko zonose. Rezultatas gautas remiantis L. Saulio bendrąja Lema 1 [4], apjungiant charakteristinių funkcijų ir kumuliantų metodus. Darbas atskiru atveju praplečia at.d. sumavimo rezultatus gautus S.A. Book [7].